

# SPARCC: Semi-Parametric Robust Estimation in a Right-Censored Covariate Model

Brian Richardson

Department of Statistical Sciences, Wake Forest University

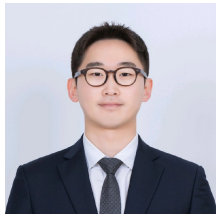
August 09, 2026



Scan for slides

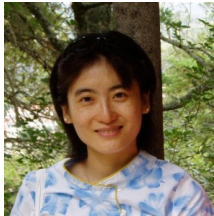
# Acknowledgements

Seong-Ho Lee, PhD



University of Seoul

Yanyuan Ma, PhD



Pennsylvania State  
University

Karen Marder, MD



Columbia University  
Medical Center

Tanya Garcia, PhD



UNC Chapel Hill

This research was supported by the National Institutes of Neurological Disorders and Stroke (grant R01NS131225) and of Environmental Health Sciences (grant T32ES007018)

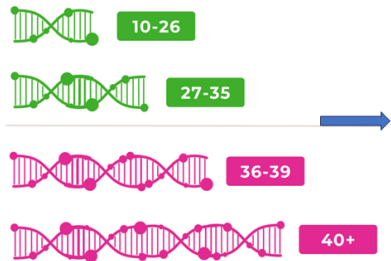
# Huntington's Disease



# Huntington's Disease

## Cause

Mutation: extra C-A-G repeats



# Huntington's Disease

## Cause

Mutation: extra C-A-G repeats



10-26



27-35



36-39



40+



## Symptoms



Cognitive

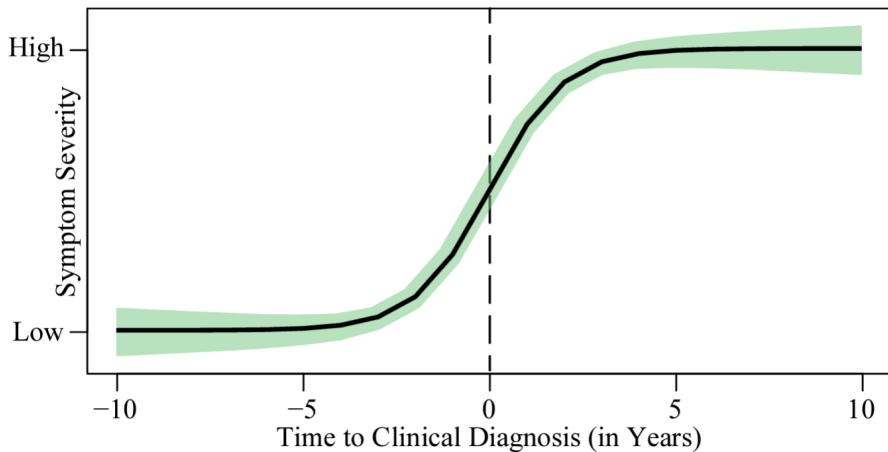


Motor



Functional

# Huntington's Disease



(Lotspeich et. al., 2024)

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **problem:**  $X$  censored by  $C$

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **problem:**  $X$  censored by  $C$ 
  - observe  $W = \min(X, C)$

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **problem:**  $X$  censored by  $C$ 
  - observe  $W = \min(X, C)$
  - observe  $\Delta = I(X \leq C)$

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **problem:**  $X$  censored by  $C$ 
  - observe  $W = \min(X, C)$
  - observe  $\Delta = I(X \leq C)$
- **nuisance:** distributions  $(f_X, f_C) = \eta$

# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **problem:**  $X$  censored by  $C$ 
  - observe  $W = \min(X, C)$
  - observe  $\Delta = I(X \leq C)$
- **nuisance:** distributions  $(f_X, f_C) = \eta$

“Humans are precious”

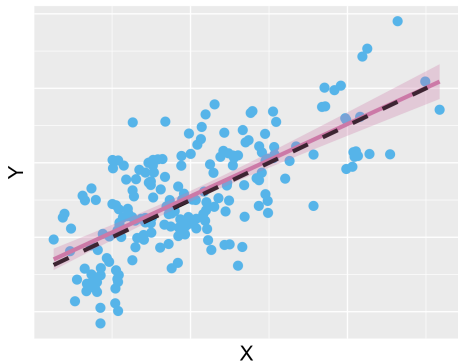
# Censoring in Huntington's Disease Studies

- **model:**  $E(Y|X) = m(X, \beta)$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **problem:**  $X$  censored by  $C$ 
  - observe  $W = \min(X, C)$
  - observe  $\Delta = I(X \leq C)$
- **nuisance:** distributions  $(f_X, f_C) = \eta$

“Humans are precious”

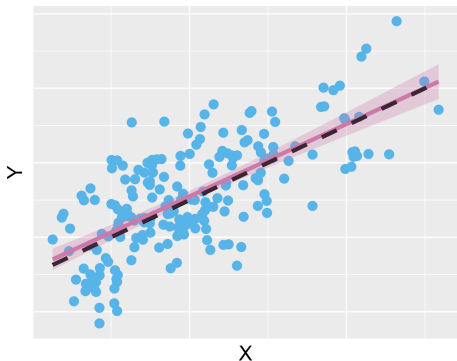
Want methods that are **robust** and **efficient**

# Censored Covariates: a Simple Example



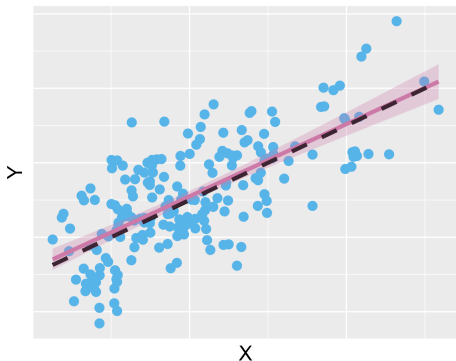
- Regression model:  $E(Y|X) = \beta_0 + \beta_1 X$

# Censored Covariates: a Simple Example



- Regression model:  $E(Y|X) = \beta_0 + \beta_1 X$
- Estimate  $\beta = (\beta_0, \beta_1)^T$  with least squares/maximum likelihood

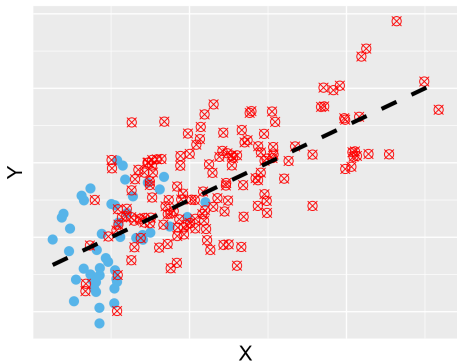
# Censored Covariates: a Simple Example



- Regression model:  $E(Y|X) = \beta_0 + \beta_1 X$
- Estimate  $\beta = (\beta_0, \beta_1)^T$  with least squares/maximum likelihood
- Solve estimating equation

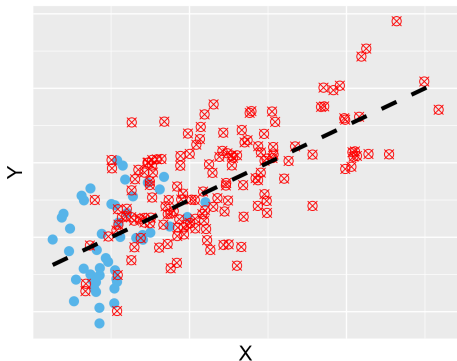
$$\sum_{i=1}^n (Y_i - \beta_0 - \beta_1 X_i)(1, X_i)^T = \mathbf{0}$$

# Censored Covariates: a Simple Example



Problem:  $X$  is censored by a **random censoring time  $C$**

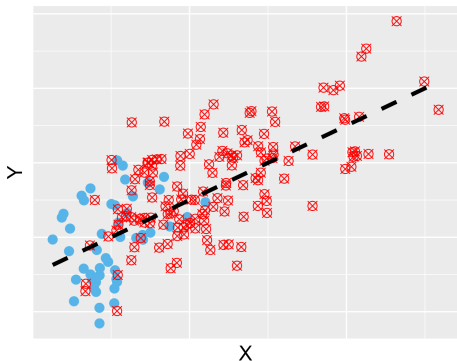
# Censored Covariates: a Simple Example



Problem:  $X$  is censored by a **random censoring time**  $C$

- $W = \min(X, C)$
- $\Delta = I(X \leq C)$

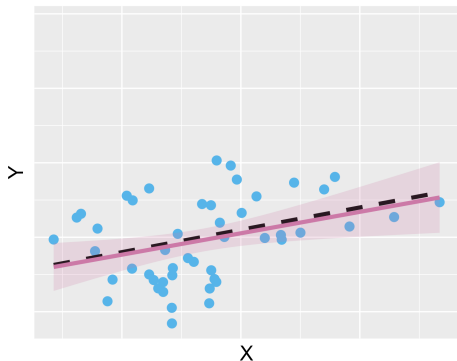
# Censored Covariates: a Simple Example



Problem:  $X$  is censored by a **random censoring time**  $C$

- $W = \min(X, C)$
- $\Delta = I(X \leq C)$
- assume:  $C \perp\!\!\!\perp (X, Y)$

# Complete Case Analysis

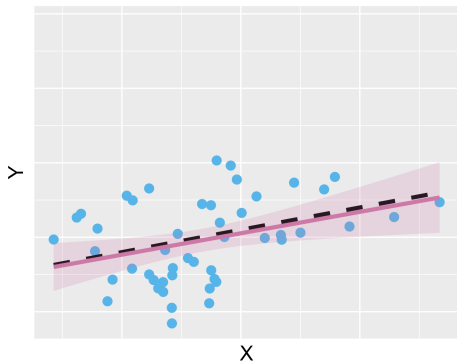


Only use *uncensored* observations

Solve estimating equation

$$\sum_{i=1}^n \Delta_i (Y_i - \beta_0 - \beta_1 W_i) (1, W_i)^T = \mathbf{0}$$

# Complete Case Analysis



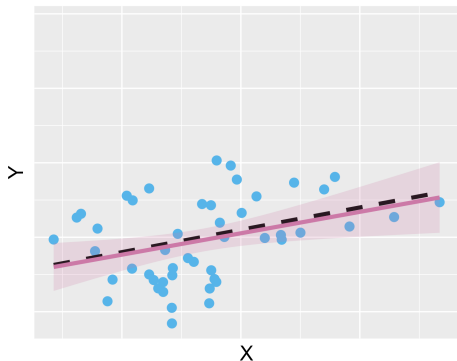
Only use *uncensored* observations

Solve estimating equation

$$\sum_{i=1}^n \Delta_i (Y_i - \beta_0 - \beta_1 W_i) (1, W_i)^T = \mathbf{0}$$

✓ Consistent

# Complete Case Analysis



Only use *uncensored* observations

Solve estimating equation

$$\sum_{i=1}^n \Delta_i (Y_i - \beta_0 - \beta_1 W_i)(1, W_i)^T = \mathbf{0}$$

✓ Consistent

✗ Inefficient

# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha) \propto \underbrace{\{f_{Y|X}(y, w, \beta)\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \beta) f_X(x, \alpha) dx \right\}^{1-\delta}}_{\text{censored}}$$

# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha) \propto \underbrace{\{f_{Y|X}(y, w, \beta)\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \beta) f_X(x, \alpha) dx \right\}^{1-\delta}}_{\text{censored}}$$

$$\mathbf{S}_{\text{ML}}(y, w, \delta, \beta) \equiv \frac{\partial}{\partial \beta} \log f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha), \quad \sum_{i=1}^n \mathbf{S}_{\text{ML}}(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha) \propto \underbrace{\{f_{Y|X}(y, w, \beta)\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \beta) f_X(x, \alpha) dx \right\}^{1-\delta}}_{\text{censored}}$$

$$\mathbf{S}_{\text{ML}}(y, w, \delta, \beta) \equiv \frac{\partial}{\partial \beta} \log f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha), \quad \sum_{i=1}^n \mathbf{S}_{\text{ML}}(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

- ✓ consistent
- ✓ fully efficient

# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha) \propto \underbrace{\{f_{Y|X}(y, w, \beta)\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \beta) f_X(x, \alpha) dx \right\}^{1-\delta}}_{\text{censored}}$$

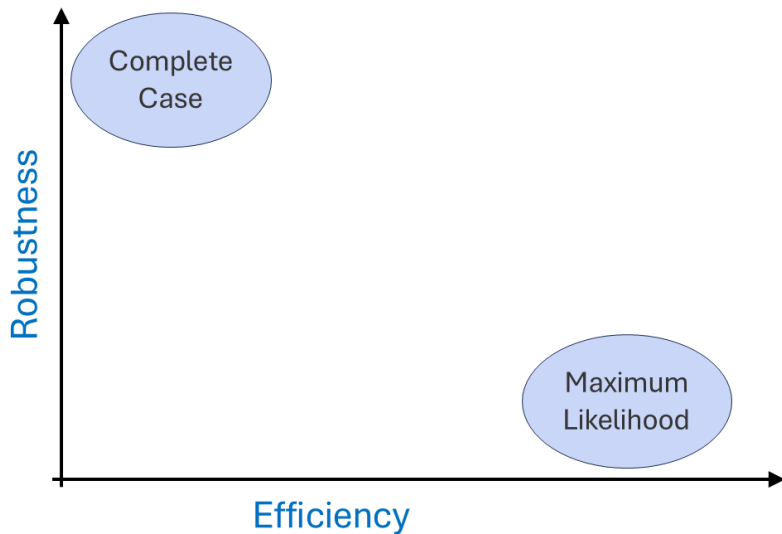
$$\mathbf{S}_{\text{ML}}(y, w, \delta, \beta) \equiv \frac{\partial}{\partial \beta} \log f_{Y,W,\Delta}(y, w, \delta, \beta, \alpha), \quad \sum_{i=1}^n \mathbf{S}_{\text{ML}}(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

✓ consistent

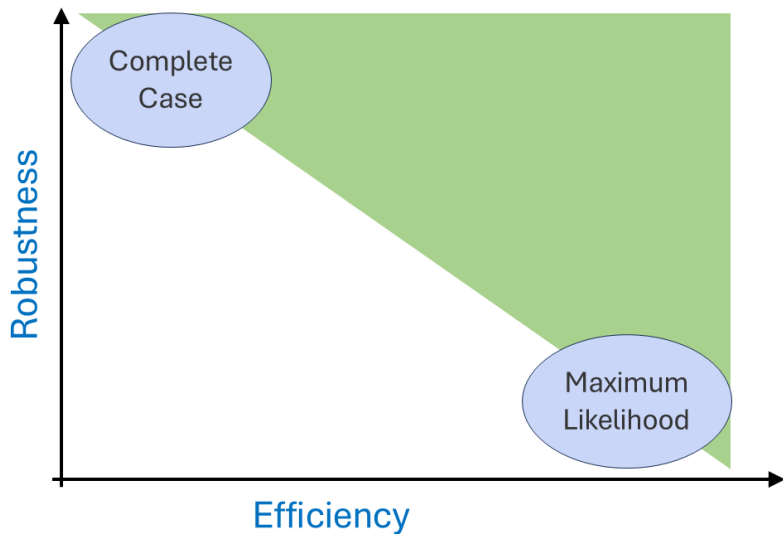
✗ inconsistent when model for  $f_X$  is incorrect

✓ fully efficient

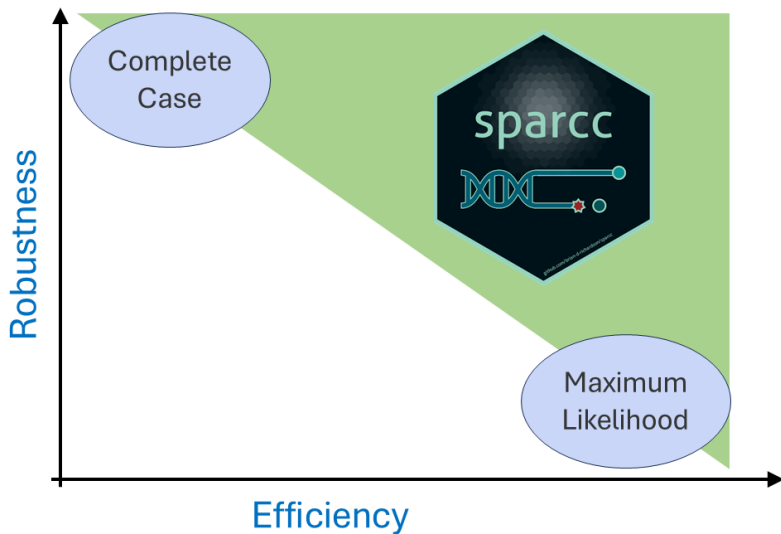
## Existing Opportunity



## Existing Opportunity



## Existing Opportunity



# The Semiparametric Recipe

- **model:**  $E(Y|X) = \beta_0 + \beta_1 X$

# The Semiparametric Recipe

- **model:**  $E(Y|X) = \beta_0 + \beta_1 X$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$

# The Semiparametric Recipe

- **model:**  $E(Y|X) = \beta_0 + \beta_1 X$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **nuisance:** distributions  $(f_X, f_C) = \eta$

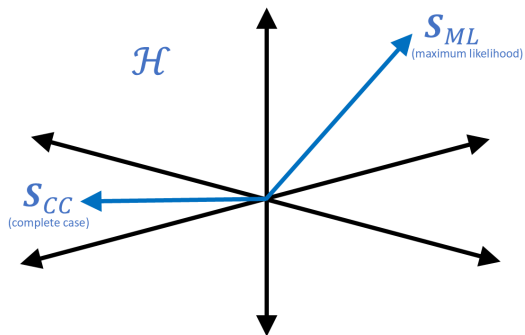
# The Semiparametric Recipe

- **model:**  $E(Y|X) = \beta_0 + \beta_1 X$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **nuisance:** distributions  $(f_X, f_C) = \eta$
- **semiparametric:** avoid parametric assumptions for  $\eta$

# The Semiparametric Recipe

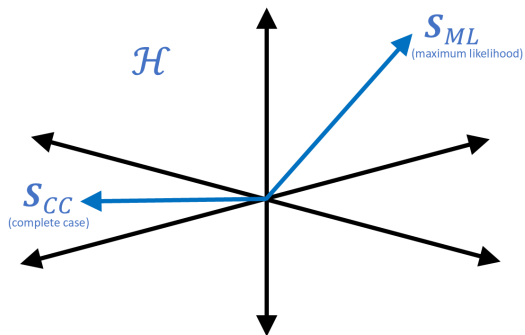
- **model:**  $E(Y|X) = \beta_0 + \beta_1 X$
- **of interest:** parameter  $\beta$  characterizing  $Y|X$
- **nuisance:** distributions  $(f_X, f_C) = \eta$
- **semiparametric:** avoid parametric assumptions for  $\eta$
- **goal:** derive the SPARCC estimator

# The Semiparametric Recipe



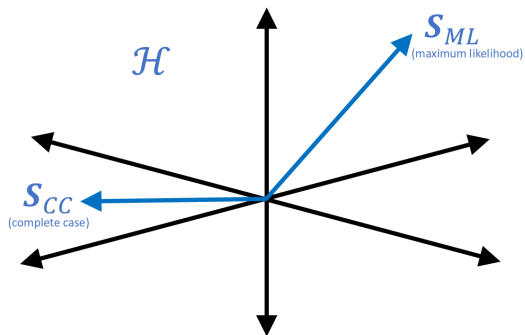
- **Hilbert space** of estimating functions

# The Semiparametric Recipe



- **Hilbert space** of estimating functions
- **covariance inner product**  $\langle \mathbf{h}, \mathbf{g} \rangle \equiv E(\mathbf{h}^T \mathbf{g})$

# The Semiparametric Recipe

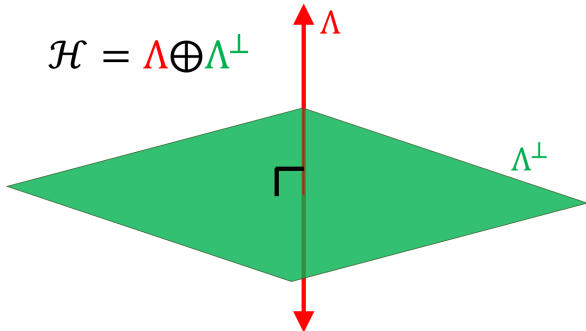


- **Hilbert space** of estimating functions
- **covariance inner product**  $\langle \mathbf{h}, \mathbf{g} \rangle \equiv E(\mathbf{h}^T \mathbf{g})$
- **orthogonal**  $\Leftrightarrow$  **uncorrelated**

$$\mathbf{h} \perp \mathbf{g} \Leftrightarrow \langle \mathbf{h}, \mathbf{g} \rangle = 0$$

# The Semiparametric Recipe

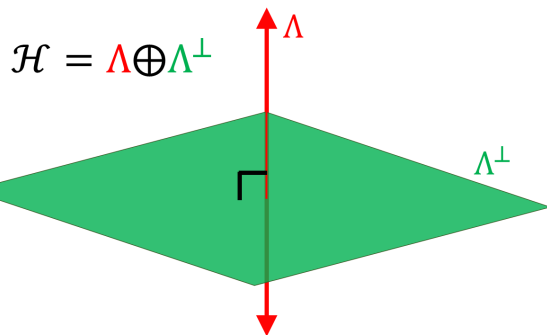
$$\mathcal{H} = \Lambda \oplus \Lambda^\perp$$



- construct  $\Lambda$  using **nuisance scores**

$$\partial \log f_{Y,W,\Delta}(y, w, \delta, \beta, \eta) / \partial \eta$$

# The Semiparametric Recipe

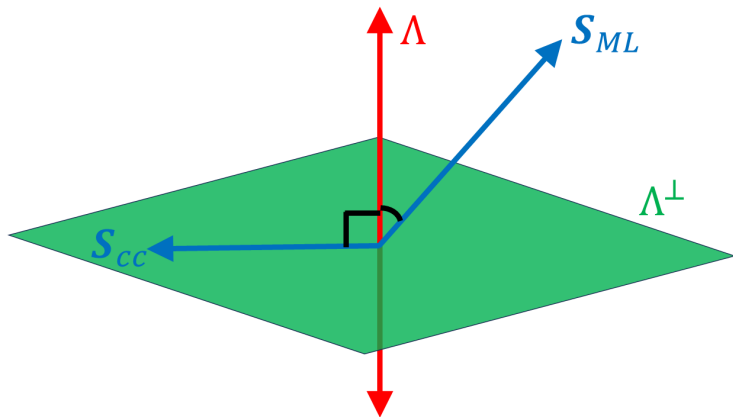


- construct  $\Lambda$  using **nuisance scores**

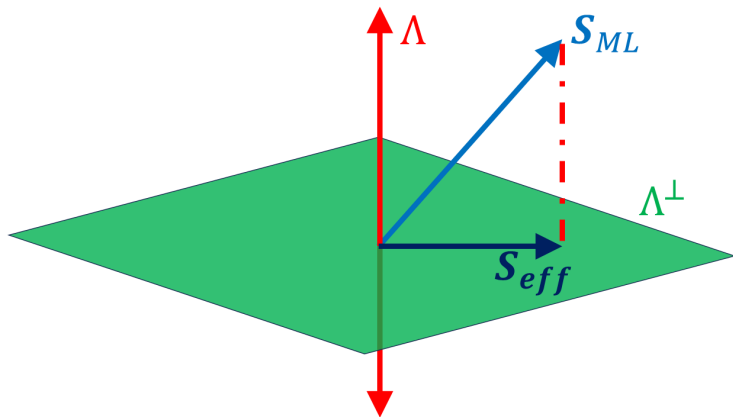
$$\partial \log f_{Y,W,\Delta}(y, w, \delta, \beta, \eta) / \partial \eta$$

- orthogonal complement  $\Lambda^\perp$

# The Semiparametric Recipe



# The Semiparametric Recipe



# Implementing the SPARCC Estimator

The **SPARCC Estimator**  $\hat{\beta}$  is the solution to

$$\sum_{i=1}^n \mathbf{s}_{\text{eff}}(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

# Implementing the SPARCC Estimator

The **SPARCC Estimator**  $\hat{\beta}$  is the solution to

$$\sum_{i=1}^n \mathbf{s}_{\text{eff}}(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

$\mathbf{S}_{\text{eff}}$  requires  $\eta = (f_X, f_C)$

# Implementing the SPARCC Estimator

The **SPARCC Estimator**  $\hat{\beta}$  is the solution to

$$\sum_{i=1}^n \mathbf{S}_{\text{eff}}(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

$\mathbf{S}_{\text{eff}}$  requires  $\eta = (f_X, f_C)$

Can be modeled either **parametrically** or **nonparametrically**

# Properties of the SPARCC Estimator

With **parametric**  $f_X, f_C, \hat{\beta}$  is:

- ✓ **doubly robust:** consistent if one of  $f_X, f_C$  is correctly specified

# Properties of the SPARCC Estimator

With **parametric**  $f_X, f_C, \hat{\beta}$  is:

- ✓ **doubly robust**: consistent if one of  $f_X, f_C$  is correctly specified
- ✓ **semiparametric efficient** if  $f_X, f_C$  are *both* correctly specified

# Properties of the SPARCC Estimator

With **parametric**  $f_X, f_C, \hat{\beta}$  is:

- ✓ **doubly robust**: consistent if one of  $f_X, f_C$  is correctly specified
- ✓ **semiparametric efficient** if  $f_X, f_C$  are *both* correctly specified

With **nonparametric**  $f_X, f_C, \hat{\beta}$  is:

- ✓ **consistent**
- ✓ **semiparametric efficient**

# Simulation Setup

$$\underbrace{Y}_{\text{outcome}} | X, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

# Simulation Setup

$$Y | \underbrace{X}_{\text{censored}}, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

# Simulation Setup

$$Y|X, \underbrace{Z}_{\text{uncensored}} \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

# Simulation Setup

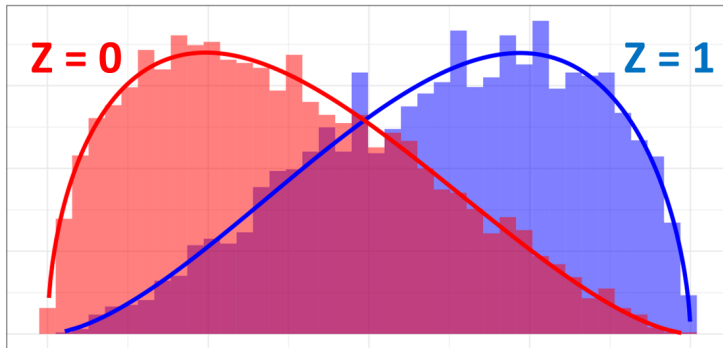
$$Y|X, Z \sim N(\underbrace{\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2}_{\text{parameter of interest: } (\beta_0, \beta_1, \beta_2, \sigma^2)})$$

# Simulation Setup: Nuisance Distributions

$$\boldsymbol{\eta} = (f_{X|Z}, f_{C|Z})$$

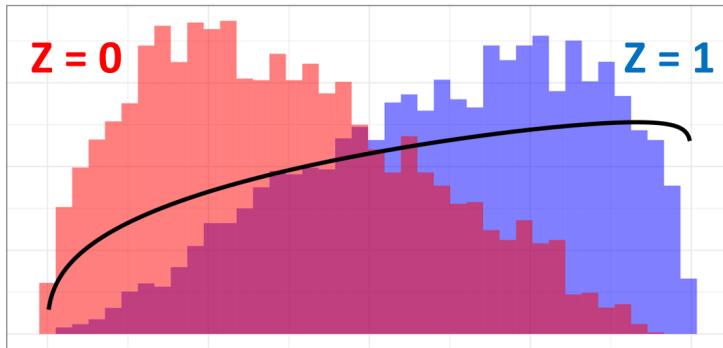
# Simulation Setup: Nuisance Distributions

$f_{X|Z}$  correct parametric:



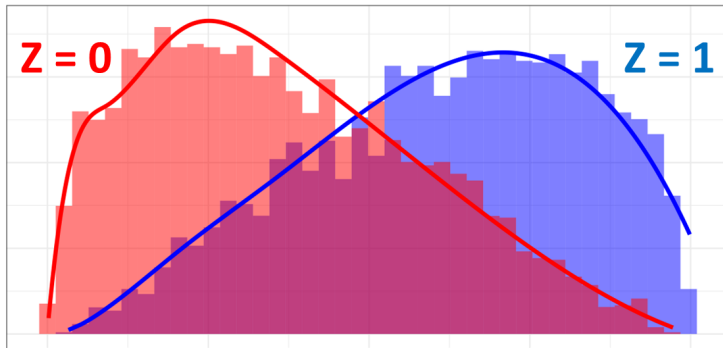
# Simulation Setup: Nuisance Distributions

$f_{X|Z}$  incorrect parametric:

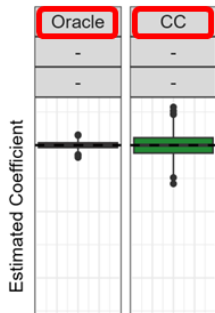


# Simulation Setup: Nuisance Distributions

$f_{X|Z}$  nonparametric:

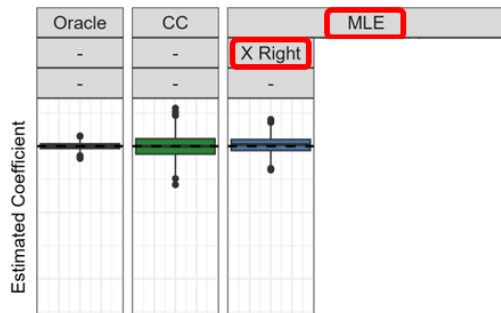


# Simulation Results: Robustness



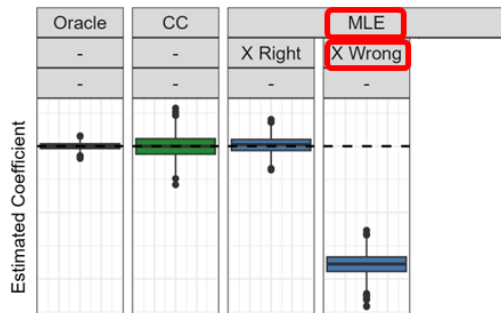
(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness



(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness



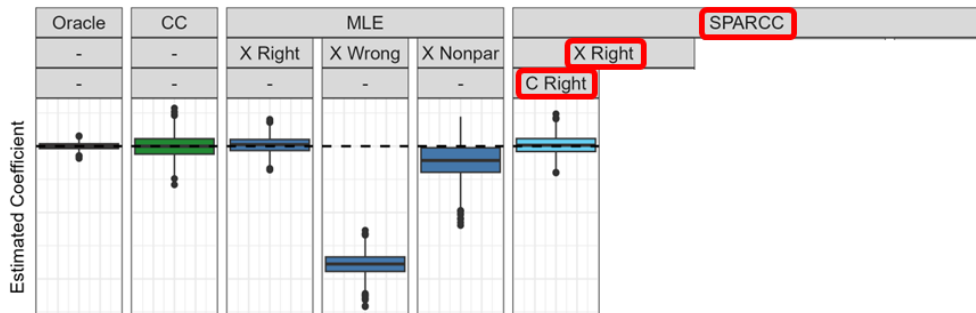
(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness



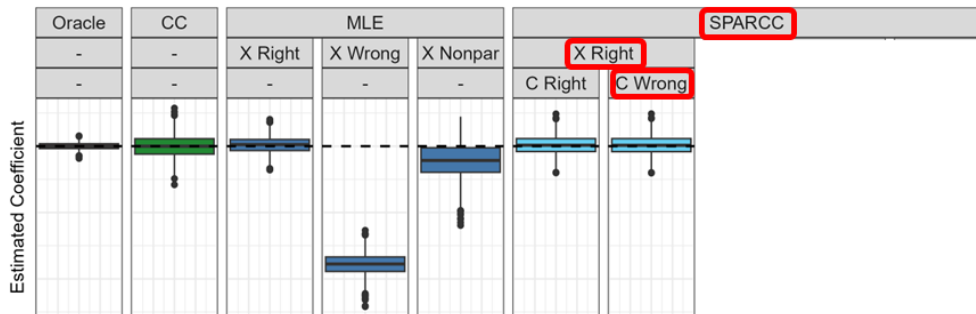
(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness



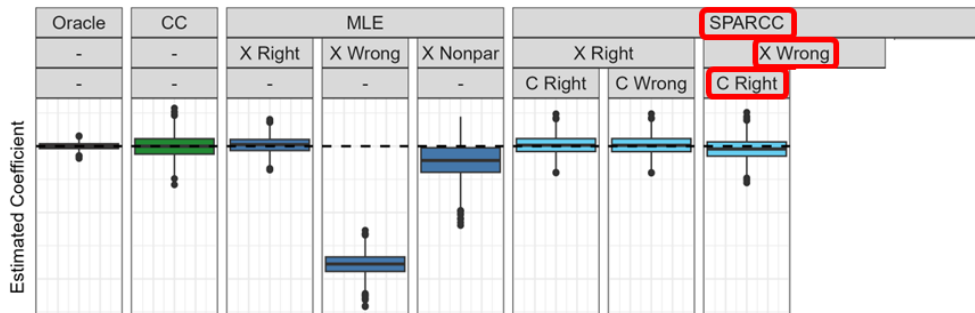
(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness



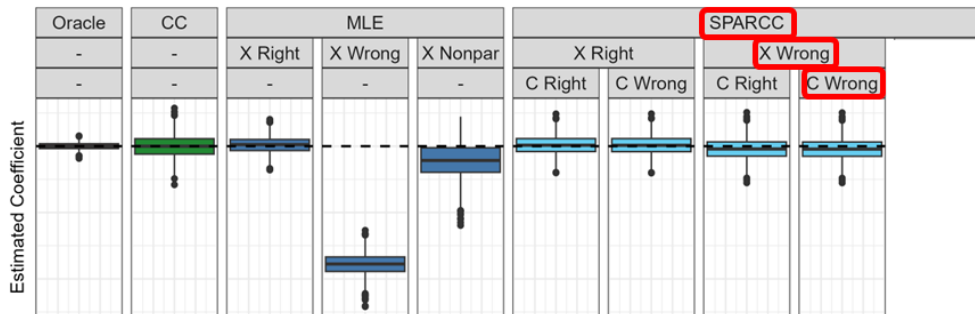
(censoring proportion  $q = 0.8$ )

## Simulation Results: Robustness



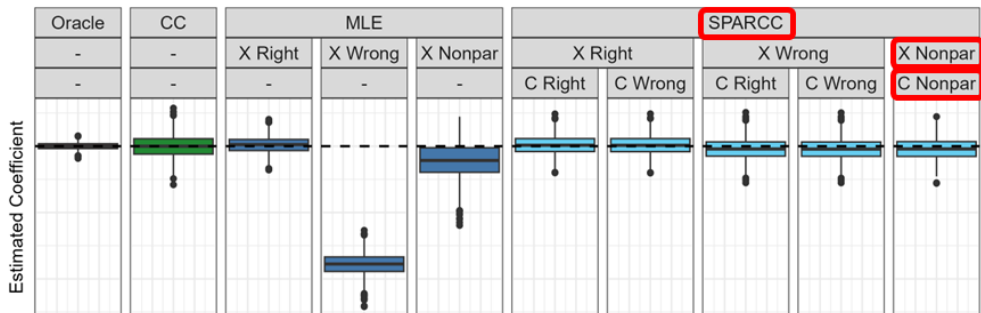
(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness



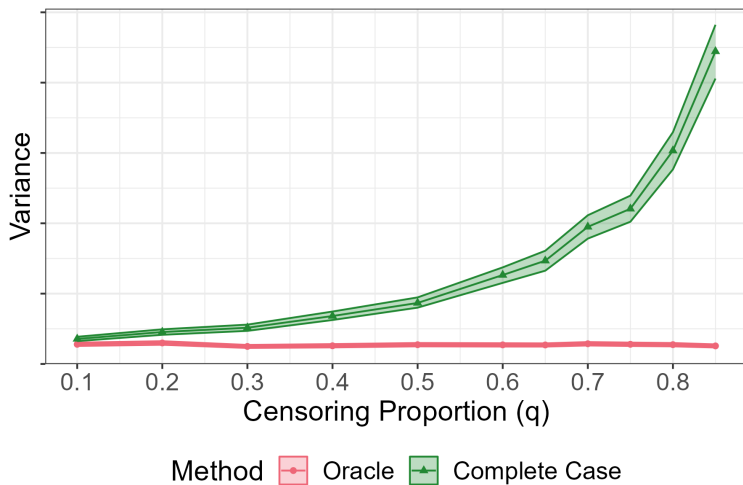
(censoring proportion  $q = 0.8$ )

# Simulation Results: Robustness

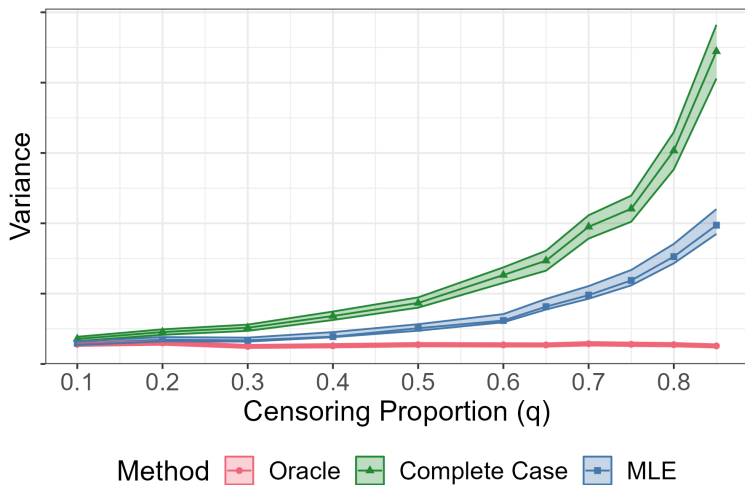


(censoring proportion  $q = 0.8$ )

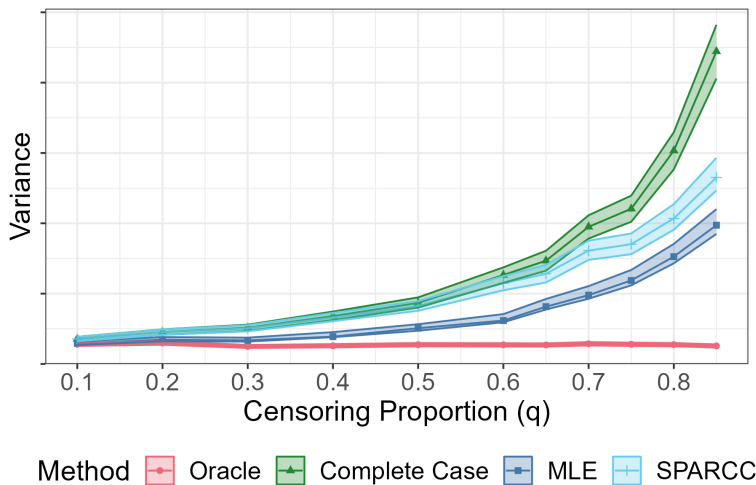
# Simulation Results: Efficiency



# Simulation Results: Efficiency



# Simulation Results: Efficiency

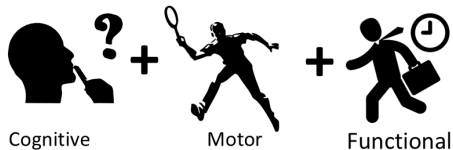


# ENROLL-HD Study

- large, observational study of people with Huntington's disease or mutation

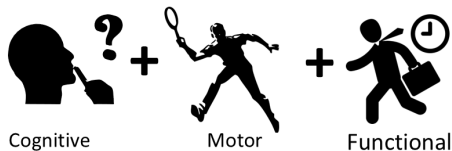
# ENROLL-HD Study

- large, observational study of people with Huntington's disease or mutation
- **composite Unified Huntington Disease Rating Scale (cUHDRS) score**



# ENROLL-HD Study

- large, observational study of people with Huntington's disease or mutation
- **composite Unified Huntington Disease Rating Scale (cUHDRS) score**



- **CAP** Score: product of age and mutation severity

# Huntington's Disease Application

$$\underbrace{Y}_{\text{cUHDS}} \mid X, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

# Huntington's Disease Application

$$Y | \underbrace{X}_{\text{time to diagnosis}}, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

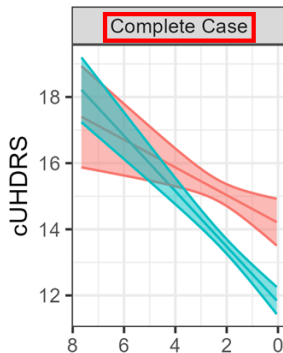
# Huntington's Disease Application

$$Y|X, \underbrace{Z}_{\text{CAP group}} \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

- **sample size:**  $n = 4530$
- **censoring rate:**  $q = 81.9\%$

# Huntington's Disease Results

Estimated Mean cUHDRS (and 95% Confidence Intervals)

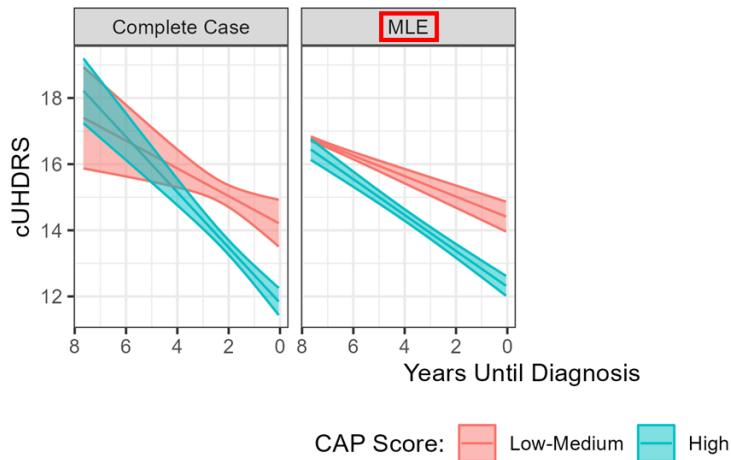


Years Until Diagnosis

CAP Score: ■ Low-Medium ■ High

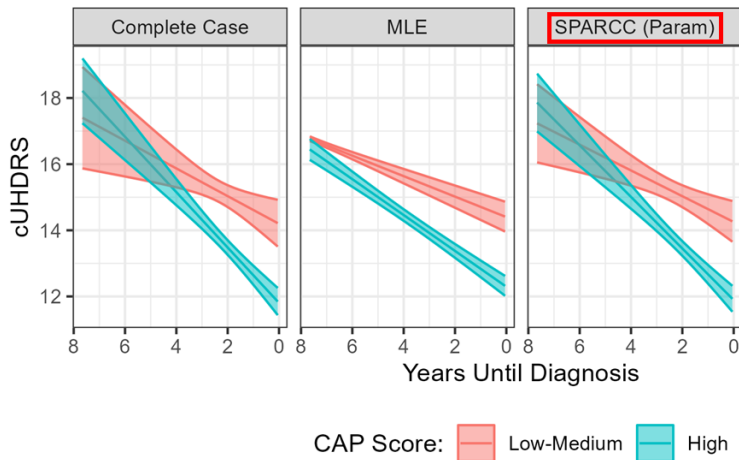
# Huntington's Disease Results

Estimated Mean cUHDRS (and 95% Confidence Intervals)



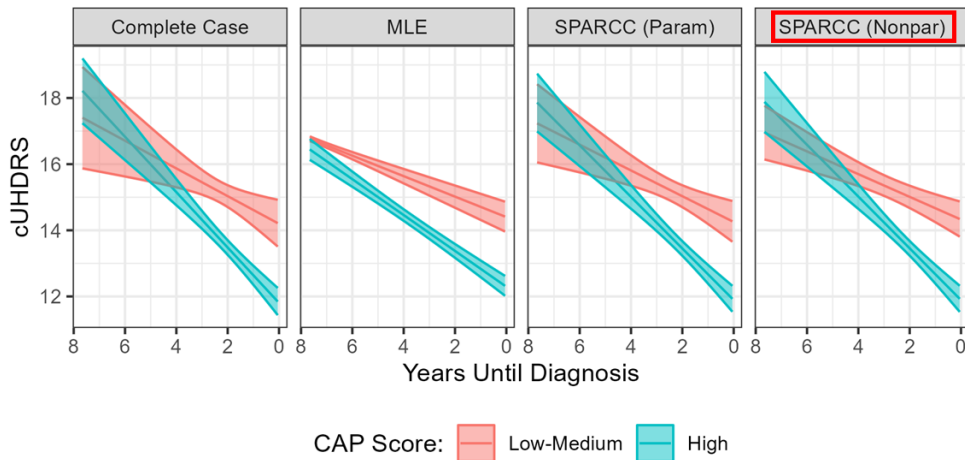
# Huntington's Disease Results

Estimated Mean cUHDRS (and 95% Confidence Intervals)



# Huntington's Disease Results

Estimated Mean cUHDRS (and 95% Confidence Intervals)



# Discussion

**Robustness** and **efficiency** matter when studying humans

# Discussion

**Robustness** and **efficiency** matter when studying humans

Proposed **SPARCC** estimator has these properties

# Discussion

**Robustness** and **efficiency** matter when studying humans

Proposed **SPARCC** estimator has these properties

## Future Work:

- Implementation and computation for general models

# Discussion

**Robustness** and **efficiency** matter when studying humans

Proposed **SPARCC** estimator has these properties

## Future Work:

- Implementation and computation for general models
- Longitudinal extension

# Discussion

**Robustness** and **efficiency** matter when studying humans

Proposed **SPARCC** estimator has these properties

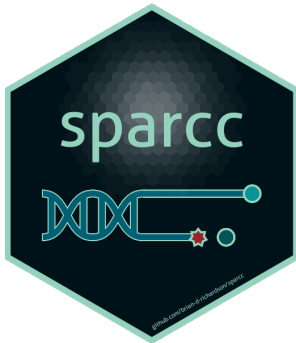
## Future Work:

- Implementation and computation for general models
- Longitudinal extension
- Allow **dependent censoring**:  $X \not\perp C|Z$

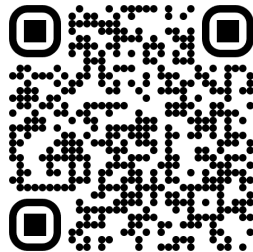
# SPARCC: **S**emi-**P**arametric Robust Estimation in a Right-**C**ensored Covariate Model



Paper in *JASA T&M*



contact: [richabr@wfu.edu](mailto:richabr@wfu.edu)



GitHub R package